

Identifying group contributions in NBA lineups with spectral analysis

Stephen Devlin* and David Uminsky

Department of Mathematics and Statistics, University of San Francisco

June 26, 2020

Abstract

We address the question of how to quantify the contributions of groups of players to team success. Our approach is based on spectral analysis, a technique from algebraic signal processing, which has several appealing features. First, our analysis decomposes the team success signal into components that are naturally understood as the contributions of player groups of a given size: individuals, pairs, triples, fours, and full five-player lineups. Secondly, the decomposition is orthogonal so that contributions of a player group can be thought of as pure: Contributions attributed to a group of three, for example, have been separated from the lower-order contributions of constituent pairs and individuals. We present detailed a spectral analysis using NBA play-by-play data and show how this can be a practical tool in understanding lineup composition and utilization.

Keywords— Basketball, lineups, group contributions, spectral analysis, representation theory

*Corresponding author, smdevlin@usfca.edu

1 Introduction

A fundamental challenge in basketball performance evaluation is the team nature of the game. Contributions to team success occur in the context of a five-player lineup, and isolating the specific contribution of an individual is a difficult problem with a considerable history. Among the many approaches to the player evaluation problem are well-known metrics like player efficiency rating (PER), wins produced (WP), adjusted plus-minus (APM), box plus-minus (BPM), win shares (WS), value over replacement player (VORP), and offensive and defensive ratings (OR and DR) to name only a few (Basketball-Reference). While these individual player metrics help create a more complete understanding of player value, some contributions remain elusive. Setting good screens, ability to draw defenders, individual defense, and off-ball movement are all examples of important contributions that are difficult to measure and quantify. In part, these contributions are elusive because they often facilitate the success of a teammate who ultimately reaps the statistical benefit.

Even beyond contributions that are difficult to quantify, the broader question of chemistry between players is a critical aspect of team success or failure. It is widely accepted that some groups of players work better together than others, creating synergistic lineups that transcend the sum of their individual parts. Indeed, finding (or fostering) these synergistic groups of players is fundamental to the role of a general manager or coach. There are, however, far fewer analytic approaches to identifying and quantifying these synergies between players. Such positive or negative effects among teammates represent an important, but much less well understood, aspect of team basketball.

In this paper we propose spectral analysis (Diaconis, 1988) as a novel approach to identifying and quantifying group effects in NBA play-by-play data. Spectral analysis is based on algebraic signal processing, a methodology that has garnered increasing attention from the machine learning community (Kakarala, 2011; Kondor et al., 2007; Kondor and Dempsey, 2012), and is particularly well suited to take advantage of the underlying structure of basketball data. The methodology can be understood as a generalization of traditional Fourier analysis, an approach whose centrality in a host of scientific and applied data analysis problems is well-known, and speaks to the promise of its application in new contexts from social choice to genetic epistasis and more (Paudel et al., 2013; Jurman et al., 2008; Lawson et al., 2006; Uminsky et al., 2018; Uminsky et al., 2019). The premise of spectral analysis in a basketball context is simple: team success (appropriately measured) can be understood as a function on lineups. Such functions have rich structure which can be analyzed and exploited for data analytic insights.

Previous work in basketball analytics has addressed similar questions from a different perspective. Both Kuehn (2016) and Maymin et al. (2013) studied lineup synergies on the level of player skills. In Maymin et al. (2013) the authors used a probabilistic framework for

game events, along with simulated games to evaluate full-lineup synergies and find trades that could benefit both teams by creating a better fit on both sides. In Kuehn (2016), on the other hand, the author used a probabilistic model to determine complementary skill categories that suggest the effect of a player in the context of a specific lineup. Work in Grassetto et al. (2019a) and Grassetto et al. (2019b) modeled lineup and player effects in the Italian Basketball League (Serie A1) based on an adjusted plus-minus framework.

Our approach is different in several respects. First, we study synergies on the level of specific player groups independent of particular skill sets. We also ignore individual production statistics and infer synergies directly from observed team success, as defined below. As a consequence of this approach, our analysis is roster constrained—we don’t suggest trades based on prospective synergies across teams. We can, however, suggest groupings of players that allow for more optimal lineups within the context of available players, a central problem in the course of an NBA game or season. Further, our approach uses orthogonality to distinguish between the contributions of a group and nested subgroups. So, for example, a group of three players that appears to exhibit positive synergies may, in fact, be benefiting from strong individual and pair contributions while the triple of players adds no particular value as a pure triple. We tease apart these higher-order correlations.

Furthermore, spectral analysis is not a model-based approach. As such, our methodology is notably free of modeling assumptions—rather than fitting the data, spectral analysis reports the observed data, albeit projected into a new basis with new information. Thus, it is a direct translation of what actually happened on the court (as we make precise below). As such, our methodology is at least complementary to existing work, and is also promising in presenting a new approach to understanding and appreciating the nuances of team basketball.

Finally, we note that while the methodology that underlies the spectral analysis approach is challenging, the resulting intuitions and insights are readily approachable. In what follows, we have stripped the mathematical details to a minimum and relegated them to references for the interested reader. The analysis, on the other hand, shows promise as a new and practical approach to a difficult problem in basketball analytics.

2 Data

We start with lineup level play-by-play data from the 2015-2016 NBA season. Such play-by-play data is publically available on ESPN.com or NBA.com, or can be purchased from websites like bigdataball.com, already processed into csv format. For a given team, we restrict attention to the 15 players on the roster having the most possessions played on the season, and filter the play-by-play data to periods of games involving only those players. Next, we compute the aggregated raw plus-minus (PM) for each lineup. Suppose lineup L plays

against opposing lineup M during a period of gameplay with no substitutions. We compute the points scored by each lineup, as well as the number of possessions for both lineups during that stretch of play. For example, if lineup L scored 6 points in 3 possessions and lineup M scored 3 points in 2 possessions, then their plus-minus is computed as the difference in points-per-possession times possessions. Thus, for L the plus-minus is $(\frac{6}{3} - \frac{3}{2})3 = 1.5$ while for M the plus-minus is $(\frac{3}{2} - \frac{6}{3})2 = -1$. Summing over all of lineup L 's possessions gives the total aggregate plus-minus for lineup L which we denote by pm_L .

Since a lineup consists of 5 players on the floor, there are $3003 = \binom{15}{5}$ possible lineups, though most see little or no playing time. We thus naturally arrive at a function on lineups by associating with L the value of that lineup's aggregate plus-minus, and write $f(L) = \text{pm}_L$. We call f the team success function. This particular success metric has the advantage of being simple and intuitive. Moreover, by summing over all lineups we recover the value of the team's cumulative plus-minus, which is highly correlated with winning percentage. The function f will serve as the foundation for our analysis, but we note that for what follows, any quantitative measure of a lineup's success could be substituted in its place.

3 Methodology

Our goal is now to decompose the function f in a way that sheds light on the various group contributions to team success. The groups of interest are generalized lineups, meaning groups of all sizes, from individual players to pairs, triples, groups of four, and full five-player lineups. Our primary tool is spectral analysis, which uses the language of representation theory (Serre, 2012) to understand functions on lineups.

Observe that a full lineup is an unordered set of five players. Any reshuffling of the five players on the floor, or the ten on the bench, does not change the lineup under consideration. Moreover, given a particular lineup, a permutation (or reshuffling) of the fifteen players on the team will result in a new lineup. The set of such permutations has a rich structure as a mathematical group. In this case, all possible permutations of fifteen players are described by S_{15} : the symmetric group on 15 items (Dummit and Foote, 2004). Furthermore, the set X of five-player lineups naturally reflects this group structure (as a homogeneous space). Most importantly for our purposes, the set of functions on lineups has robust structure with respect to the natural action of permutations on functions. This structure is well understood and can be exploited for data analytic insights as we show below. By way of analogy, just as traditional Fourier analysis looks to decompose a time series into periodicities that can reveal a hidden structure (weekly or seasonal trends, say), our decomposition of f will reveal group effects in lineup-level data.

Let $L(X)$ denote the collection of all real valued functions on five-player lineups. This

set is a vector space with the usual notions of sum of functions, multiplication by scalars, and an inner product given by

$$\langle g, h \rangle = \frac{1}{|X|} \sum_{x \in X} g(x)h(x). \quad (1)$$

The dimension of $L(X)$ is equal to the number of lineups, $3003 = \binom{15}{5}$. In light of the permutation group's action on $L(X)$ as mentioned above, $L(X)$ admits a natural (invariant and irreducible) decomposition as follows:

$$L(X) = V_0 \oplus V_1 \oplus V_2 \oplus V_3 \oplus V_4 \oplus V_5. \quad (2)$$

Each V_i , with $0 \leq i \leq 5$ is a vector subspace with data analytic significance. Rather than give a self contained treatment of this decomposition, we refer to Diaconis (1988) and Dummit and Foote (2004), and here, simply note that each space is spanned by the matrix coefficients of the irreducible representations of the group S_{15} associated with Young tableaux of shape $(10, 5)$. We can gain some intuition for the decomposition by considering the lower-order spaces as follows. An explicit computation of the decomposition is given in section 4 below for a toy example.

Take δ_L to be the indicator function of a fixed lineup L , so that $\delta_L(L) = 1$, while $\delta_L(L') = 0$ for any other lineup L' . As above, X is the set of all possible lineups, and

$$\delta = \sum_{L \in X} \delta_L. \quad (3)$$

If we act on the function δ by reshuffling lineups (this is the action of the permutation group S_{15}), we see that while the terms in the summation in (3) get reordered, the function itself remains unchanged. (See section 4 below for details.) Thus, the one-dimensional space spanned by δ is invariant under lineup reshuffling and represents the mean value of the function f since we can write $f = c\delta + (f - c\delta)$. Here, c is just the average value of f and $c\delta$ is the best possible constant approximation to f . The function $f - c\delta$ represents the original data, but now centered with mean zero, and orthogonal to the space of constant functions with respect to the inner product in (1). The space spanned by δ is V_0 in (2).

To understand V_1 , we start with indicator functions for individual players. Given a player i , define $\delta_i = \sum_{L \in \mathcal{L}_i} \delta_L - m\delta$ where the sum is over all lineups that include player i and four other players, and m is a constant chosen so that δ_i is orthogonal to δ . One can show that the space spanned by $\{\delta_1, \delta_2, \dots, \delta_{15}\}$ is again stable under lineup reshuffling. (Though the set of individual indicator functions is linearly dependent, and only spans a 14-dimensional space as we'll see below.)

The decomposition continues in an analogous way, though the computations become more involved. Several computational approaches are described in Diaconis (1988) and Maslen et al. (2003). In our case of the symmetric group S_{15} acting on lineups, we employ the method in Maslen et al. (2003), which involves first computing the adjacency matrix of an associated *Johnson graph* $J(15, 5)$. It turns out that $J(15, 5)$ has 6 eigenvalues, each of which is associated with one of the effect spaces: zero (mean), and first through fifth-order spaces. Specifically, the largest eigenvalue is simple and is associated with the one-dimensional mean space; the second largest eigenvalue is associated with the first-order space, etc. It is now a matter of computing an eigenbasis for each space, and using it to project the data vector onto each eigenspace to give the orthogonal decomposition used in (2). It is also worth noting that spectral analysis includes the traditional analysis of variance as a special case, a connection suggested by the discussion above and further explained in Diaconis (1988).

The decomposition in (2) is particularly useful for two reasons. First, each V_i can be interpreted as the space of functions encoding i -th order effects. For instance, one can see that V_1 is naturally understood as encoding first-order individual effects beyond the mean. Thus, the projection of f onto V_1 can be understood as that part of team success f attributable to the contributions of individual players. Similarly V_2 includes effects attributable to pure player pairs (individual contributions have been removed), and the corresponding projection of f in V_2 gives the contributions of those pairs to team success. V_3 encodes contributions of groups of three, and so on. These interpretations follow from the fact that each subspace in the decomposition of $L(X)$ is invariant under the natural reshuffling action of S_{15} on lineups. It is also worth noticing that the lineup success function is completely recovered via its projections onto the order subspaces in (2). If we write f_i for the projection of f onto V_i , then $f = f_0 + f_1 + f_2 + f_3 + f_4 + f_5$. As such, the spectral decomposition gives a complete description of the original data set with respect to a new basis grounded in group contributions.

Secondly, the decomposition in (2) is orthogonal (signified by the \oplus notation). From a data analytic perspective, this means that there is no overlap among the spaces, and group effects are independent. Thus, for instance, a contribution attributed to a group of three players can be understood as a pure third-order contribution. All constituent pair and individual contributions have been removed and quantified separately in the appropriate lower-order spaces. We thus avoid erroneous attribution of success due to multicollinearity among groups. For example, is a big three really adding value as a triple, or is its success better understood as a strong pair plus an individual? The spectral decomposition in (2) provides a quantitative basis for answering such questions.

The advantage of the orthogonality of the spaces in (2), however, presents a challenge with respect to direct interpretation of contributions for particular groups. This is evident when considering the dimension of each of the respective effect spaces in Table 1, which is

Space	Dimension	Number of Groups
V_0	1	–
V_1	14	15
V_2	90	105
V_3	350	455
V_4	910	1365
V_5	1638	3003

Table 1: Dimension of each effect space, along with the number of natural groups of each size.

strictly smaller than the number of groups of that size we might wish to analyze.

Since we have rosters of fifteen players, there are fifteen individual contributions to consider. The space V_1 , however, is 14-dimensional. Similarly, while V_2 includes all of the contributions to f attributable to pairs of players, it does so in a 90-dimensional space despite the fact that there are $105 = \binom{15}{2}$ natural pairs of players to consider. The third-order space V_3 has dimension 350 while there are 455 player triples, and so on.

We deal with this issue using Mallows’ method of following easily interpretable vectors as in Diaconis (1988). Let g be a group of players. For example, if players are labeled 1 through 15, then a particular triple might be $g = \{1, 2, 7\}$. Let ϕ_g be the indicator function associated with g , i.e., the function that takes the value 1 when all three players 1, 2, and 7 are in a lineup, and outputs 0 otherwise. The function ϕ_g is intuitively associated with the success of the group g (though it is not invariant under reshuffling and is not orthogonal to nested lower-order groups).

To quantify the contribution of g (as a pure triple) to the success of the team as measured by f , project both ϕ_g and f onto V_3 and take the inner product of the projections: $\langle pr_{V_3}(\phi_g), pr_{V_3}(f) \rangle = \langle pr_{V_3}(\phi_g), f_3 \rangle$. After projecting onto V_3 we are left with only the third-order components of ϕ_g and f . The resulting inner product is a weighted cosine similarity that indicates the extent to which the pure triple g is correlated with the team’s success f . Larger values of this inner product reflect a stronger synergy between the triple of players $\{1, 2, 7\}$, while a negative value indicates that, after removing the contributions of the constituent individuals and pairs, spectral analysis finds this particular group of three ineffective. In the results below we show how this information might be useful in evaluating lineups.

4 Two-On-Two Basketball

To ground the ideas of the previous section we present a small-scale example in detail. Consider a version of basketball where a team consists of 5 players, two of which play at any given moment. The set of possible lineups consists of the ten unordered pairs $\{i, j\}$ with $i, j \in \{1, 2, 3, 4, 5\}$ and $i \neq j$. The symmetric group S_5 acts on lineups by relabeling, and we extend this action to functions on lineups as follows. Given a permutation π , a function h , and a lineup L , define

$$(\pi \cdot h)(L) = h(\pi^{-1}L). \quad (4)$$

Therefore, if π is the permutation (123), taking player 1 to player 2, player 2 to player 3, player 3 to player 1, and leaving everyone else fixed, and if L is the lineup $\{1, 3\}$, then

$$(\pi \cdot h)(L) = h(\pi^{-1}\{1, 3\}) = h(\{3, 2\}). \quad (5)$$

The use of the inverse is necessary to ensure that the action on functions respects the operation in the group, that is, so that $(\tau\pi) \cdot h = \tau \cdot (\pi \cdot h)$ (Dummit and Foote, 2004).

Following a season of play, we obtain a success function that gives the plus-minus (or other success metric) of each lineup. We might observe a function like that in Table 2.

L	$f(L)$	L	$f(L)$
$\{1, 2\}$	22	$\{2, 4\}$	35
$\{1, 3\}$	18	$\{2, 5\}$	26
$\{1, 4\}$	3	$\{3, 4\}$	84
$\{1, 5\}$	58	$\{3, 5\}$	25
$\{2, 3\}$	93	$\{4, 5\}$	2

Table 2: Success function for two-player lineups.

Summing $f(L)$ over all lineups that include a particular player gives individual raw plus-minus as in Table 3. Player 3 is the top rated individual, followed by 2, 4, 5, and 1. Lineup rankings are given by $f(L)$ itself, which shows $\{2, 3\}$, $\{3, 4\}$, and $\{1, 5\}$ as the top three.

Now compare the analysis above with spectral analysis. In this context the vector space of functions on lineups is 10-dimensional and has a basis consisting of vectors $\delta_{\{i,j\}}$ that assign the value 1 to lineup $\{i, j\}$ and 0 to all other lineups. The decomposition in (2) becomes

$$V = V_0 \oplus V_1 \oplus V_2. \quad (6)$$

Define $\delta = \sum_{\{i,j\}} \delta_{\{i,j\}}$. The span of δ is the one-dimensional subspace V_0 of constant functions. Moreover, V_0 is S_5 invariant since for any relabeling of players given by π , we

Player	PM	Rank
1	101	5
2	176	2
3	220	1
4	124	3
5	111	4

Table 3: Preliminary analysis of sample team using individual plus-minus (PM), which is the sum of the lineup PM over lineups that include a given individual.

have $\pi \cdot \delta = \delta$. Given a function f in V , its projection f_0 on V_0 will assign to each lineup the average value of f , in this case 36.6.

First order (or individual) effects beyond the mean are encoded in V_1 . Explicitly, define $\delta_1 = \sum_i \delta_{\{1,i\}} - \frac{2}{5}\delta$, with δ_2, δ_3 , and δ_4 defined analogously. One can check that the 4-dimensional vector space spanned by $\{\delta_1, \delta_2, \delta_3, \delta_4\}$, is S_5 invariant, and is orthogonal to V_0 . Since the mean has been subtracted out and accounted for in V_0 , a vector in V_1 represents a pure first order effect. Note that $\delta_5(x) = \sum_i \delta_{\{5,i\}} - \frac{2}{5}\delta$ can be written $\delta_5 = -\delta_1 - \delta_2 - \delta_3 - \delta_4$. Consequently, V_1 is 4-dimensional even though there are five natural first order effects to consider: one for each player.

Finally, the orthogonal complement of $V_0 \oplus V_1$ is the 5-dimensional S_5 invariant subspace V_2 . V_2 gives the contribution to f from *pure pairs*, or pure second order effects after the mean and individual contributions are removed. The three subspaces V_0, V_1 , and V_2 are all irreducible since none contains a nontrivial S_5 invariant subspace.

We can now project f onto V_0, V_1 , and V_2 . All together we have $f = f_0 + f_1 + f_2$:

$$f \begin{pmatrix} \{1, 2\} \\ \{1, 3\} \\ \{1, 4\} \\ \{1, 5\} \\ \{2, 3\} \\ \{2, 4\} \\ \{2, 5\} \\ \{3, 4\} \\ \{3, 5\} \\ \{4, 5\} \end{pmatrix} = \begin{bmatrix} 22 \\ 18 \\ 3 \\ 58 \\ 93 \\ 35 \\ 26 \\ 84 \\ 25 \\ 2 \end{bmatrix} = \begin{bmatrix} 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \\ 36.6 \end{bmatrix} + \begin{bmatrix} -5.27 \\ 9.40 \\ -22.60 \\ -26.93 \\ 34.40 \\ 2.40 \\ -1.93 \\ 17.07 \\ 12.73 \\ -19.27 \end{bmatrix} + \begin{bmatrix} -9.33 \\ -28.00 \\ -11.00 \\ 48.33 \\ 22.00 \\ -4.00 \\ -8.67 \\ 30.33 \\ -24.33 \\ -15.33 \end{bmatrix} \quad (7)$$

Turning to the question of interpretability, section 3 proposes Mallows' method of using readily interpretable vectors projected into the appropriate effect space. To that end, the individual indicator function $\phi_{\{2\}} = \delta_{\{1,2\}} + \delta_{\{2,3\}} + \delta_{\{2,4\}} + \delta_{\{2,5\}}$ is naturally associated with

player 2: $\phi_{\{2\}}(L) = 1$ when player 2 is in L and is 0 otherwise. We quantify the effect of player 2 by projecting $\phi_{\{2\}}$ and f into V_1 , and then taking the dot product of the projections. For a lineup like $\{2, 3\}$, we take the dot product of the projections of the lineup indicator function $\delta_{\{2,3\}}$, and f , in V_2 . Note that player 2's raw plus-minus is the inner product of $10 \cdot f$ with the interpretable function $\phi_{\{2\}}$. Similarly $f(\{i, j\})$ is $10 \cdot \langle f, \phi_{\{i,j\}} \rangle$. The key difference is that spectral analysis uses Mallows' Method *after* projecting onto the orthogonal subspaces in (6).

Contributions from spectral analysis as measured by Mallows' method are given in Table 4 for both individuals and (two-player) lineups. The table also includes both the spectral

Individual	Spec	Pair	Spec	Rank	f Rank	Pair	Spec	Rank	f Rank
{1}	-45.4	{1,2}	-9.3	6	7	{2,4}	-4	4	4
{2}	29.6	{1,3}	-28	10	8	{2,5}	-8.7	5	5
{3}	73.6	{1,4}	-11	7	9	{3,4}	30.3	2	2
{4}	-22.4	{1,5}	48.3	1	3	{3,5}	-24.3	9	6
{5}	-35.4	{2,3}	22	3	1	{4,5}	-24	8	10

Table 4: Spectral value (Spec) for each individual player and two-player lineup, and rank of each lineup, along with the preliminary rank given by f .

and preliminary (based on f) rankings of each lineup. Note that lineup $\{2, 3\}$ drops from the best pair to the third best pure pair. Once we account for the contributions of players two and three as individuals, the lineup is not nearly as strong as it appears in the preliminary analysis. We find stronger pair effects from lineups $\{1, 5\}$ and $\{3, 4\}$. All remaining lineups are essentially ineffective in that their success can be attributed to the success of the constituent individuals rather than the pairing. Interesting questions immediately arise. What aspects of player four's game result in a more effective pairing with player three, the team's star individual player, than the pairing of three with two, the team's second best individual? What is behind the success of the $\{1, 5\}$ lineup? These considerations are relevant to team construction, personnel considerations, and substitution patterns. We pursue this type of analysis further in the context of an actual NBA team below.

5 Results and Discussion

A challenge inherent in working with real lineup-level data is the wide disparity in the number of possessions that lineups play. Most teams have a dominant starting lineup that plays far more possessions than any other. For example, the starting lineup of the '16 Golden State Warriors played approximately 1140 possessions while the next most used lineup played 535 possessions. Only 12 lineups played more than 100 possessions for the Warriors on the season.

For the Boston Celtics, the starters played 1413 possessions compared to 257 for the next most utilized, with 13 lineups playing more than 100 possessions. By contrast, the Celtics had 255 lineups that played fewer than 10 possessions (but at least one), and the Warriors had 236. Numbers are similar across the league. This is another reason for using raw plus-minus in defining the team success function f on lineups. A metric like per-possession lineup plus-minus breaks down in the face of large numbers of very low possession lineups and a few high possession lineups. Still, we want to identify potentially undervalued and underutilized groups of players— especially for smaller groups like pairs and triples where there are many more groups that do play significant numbers of possessions. Another consideration is that over time, lineups with large numbers of possessions will settle closer to their true mean value while lineups with few possessions will be inherently noisier. As a result, we perform the spectral analysis on f as described in section 3 above, and then normalize the spectral contribution by the log of possessions played by each group. We call the result *spectral contribution per log possession* (SCLP). This balances the considerations above and allows strong lower possession groups to emerge while not over-penalizing groups that do play many possessions.

Despite these challenges, however, we'll see below that there are significant insights to be gained in working with lineup level data. Moreover, since spectral analysis is a non-model-based description of complete lineup-level game data, it has the advantage of maintaining close proximity to the actual gameplay observed by coaches, players, and fans. There are always five players on the floor, so all data begins at the level of full lineups.

Consider the first order effects for the 15-16 Golden State Warriors in Table 5. Draymond Green, Stephen Curry, and Klay Thompson are the top three players. The ordering, specifically Green ranked above Curry, is perhaps interesting, though it's worth noting that this ordering agrees with ESPN's real plus-minus (RPM). (Green led the entire league in RPM in 15-16.) Other metrics like box plus-minus (BPM) and wins-above-replacement (WAR) rank Curry higher. Because SCLP is based on ability of lineups to outscore opponents when the player is on the floor (like RPM), however, as opposed to metrics like BPM and WAR which are more focused on points produced, the ordering is defensible.

In fact, a closer look at the interpretable vector ϕ_i associated with individual player i (as described in sections 3 and 4) reveals that $\phi_i = \delta_i + c \cdot \delta$, so is just a non-mean-centered version of the first order invariant functions that span V_1 . Consequently, the spectral contribution (non-possession normalized) is a linear function of individual plus-minus, so reflects precisely that ordering. This is not the case for higher-order groups, however, which is where we focus the bulk of our analysis.

The second-order effects are given in in Table 6, and quantify the contributions of player pairs, having removed the mean, individual, and higher-order group effects. The top and bottom five pairs (in terms of SCLP) are presented here, with more complete data in Table

Player	SCLP	PM	Poss
Draymond Green	17.2	1038.4	5800
Stephen Curry	15.9	978.7	5610
Klay Thompson	12.0	808.6	5453
Andre Iguodala	3.5	436.1	3516
Andrew Bogut	2.8	403.6	2951
Marreese Speights	-7.4	20.0	1630
Ian Clark	-9.8	-51.9	1108
Anderson Varejao	-11.1	-34.4	368
Jason Thompson	-11.2	-33.8	339
James Michael McAdoo	-12.1	-85.0	526

Table 5: Top and bottom five first-order effects for GSW. SCLP is the spectral contribution per log possession, PM is the player’s raw plus-minus, and Poss is the number of possessions for that player.

16 in the appendix. Even after accounting for and removing their strong individual contributions, however, it is notable that Green–Curry, Curry–Thompson, and Green–Thompson are the dominant pair contributors by a considerable margin, with SCLP values that are all more than twice as large as for the next largest pair (Barbosa–Speights). These large positive SCLP values represent true synergies: These pairs contribute to team success *as pure pairs*. The fact that the individual contributions of the constituent players are also positive results in a stacking of value within a lineup that provides a quantifiable way of assessing whether the whole does indeed add to more than the sum of its parts.

Reserves Leandro Barbosa, Mareese Speights, and Ian Clark, on the other hand, were poor individual contributors, but manage to combine effectively in several pairs. In particular, the Barbosa–Speights pairing is notable as the fourth best pure pair on the team (in 983 possessions). After accounting for individual contributions, lineups that include the Barbosa–Speights pairing benefited from a real synergy that positively contributed to team success. This suggests favoring, when feasible, lineup combinations with those two players together to leverage this synergy and mitigate their individual weaknesses.

Tables 7 and 8 show pair values for players Andrew Bogut and Shaun Livingston (again in pairs with at least 150 possessions, and with more detailed tables in the appendix). Both players are interesting with respect to second order effects. While Bogut was a positive individual contributor, and was a member of the Warriors’ dominant starting lineup that season, he largely fails to find strong pairings. His best pairings are with Klay Thompson and Harrison Barnes, while he pairs particularly poorly with Andre Iguodala (in a considerable

P1	P2	SCLP	PM	Poss
Draymond Green	Stephen Curry	13.3	979.9	5102
Stephen Curry	Klay Thompson	11.2	827.8	4311
Draymond Green	Klay Thompson	11.1	847.8	4678
Leandro Barbosa	Marreese Speights	5.3	76.2	983
Draymond Green	Andre Iguodala	4.3	490.0	2165
Draymond Green	Ian Clark	-7.2	33.3	424
Klay Thompson	Leandro Barbosa	-7.2	4.8	349
Stephen Curry	Ian Clark	-8.1	14.0	220
Draymond Green	Anderson Varejao	-9.5	7.2	217
Stephen Curry	Anderson Varejao	-10.1	-26.9	237

Table 6: Top and bottom five SCLP pairs with at least 200 possessions, along with raw plus-minus and possessions.

785 possessions). This raises interesting questions as to why Bogut’s style of play is better suited to players like Thompson or Barnes rather than players like Curry or Iguodala. Also noteworthy is the fact that the Bogut–Iguodala pairing has a positive plus-minus value of 107. The spectral interpretation is that this pairing’s success should be attributed to the individual contributions of the players, and once those contributions are removed, the group lacks value as a pure pair.

P1	P2	SCLP	PM	Poss
Andrew Bogut	Klay Thompson	3.7	394.3	2637
Andrew Bogut	Harrison Barnes	2.1	206.2	1527
Andrew Bogut	Stephen Curry	1.6	378.5	2530
Andrew Bogut	Andre Iguodala	-2.1	107.0	785

Table 7: Select pairs involving Andrew Bogut (with at least 150 possessions).

P1	P2	SCLP	PM	Poss
Shaun Livingston	Anderson Varejao	2.0	-1.5	174
Shaun Livingston	Marreese Speights	1.6	17.8	1014
Shaun Livingston	Draymond Green	1.2	323.6	1486
Shaun Livingston	Andre Iguodala	-1.3	65.2	1605
Shaun Livingston	Klay Thompson	-3.6	111.8	1412

Table 8: Select pairs involving Shaun Livingston (with at least 150 possessions).

P1	P2	P3	SCLP	PM	Poss
Draymond Green	Stephen Curry	Klay Thompson	12.6	812.7	4085
Draymond Green	Klay Thompson	Harrison Barnes	5.9	427.3	2473
Draymond Green	Stephen Curry	Andre Iguodala	5.8	464.8	1830
Stephen Curry	Klay Thompson	Harrison Barnes	5.7	416.5	2431
Stephen Curry	Klay Thompson	Andrew Bogut	4.9	382.2	2296
Stephen Curry	Andre Iguodala	Brandon Rush	-3.8	-13.5	207
Draymond Green	Stephen Curry	Marreese Speights	-4.1	97.9	299
Draymond Green	Klay Thompson	Marreese Speights	-4.5	52.2	250
Draymond Green	Klay Thompson	Ian Clark	-5.8	9.8	316
Draymond Green	Stephen Curry	Ian Clark	-7.4	14.5	205

Table 9: Best and worst third-order effects for GSW with at least 200 possessions.

Shaun Livingston, on the other hand, played an important role as a reserve point guard for the Warriors. Interestingly, Livingston’s worst pairing by far was with Klay Thompson. Again, considering the particular styles of these players compels interesting questions from the perspective of analyzing team and lineup compositions and playing style. It’s also noteworthy that this particular pairing saw 1412 possessions, and it seems entirely plausible that its underlying weakness was overlooked due to the healthy 111.8 plus-minus with that pair on the floor. The success of those lineups should be attributed to other, better synergies. For example, one rotation added Livingston as a sub for Barnes (112 possessions). Another put Livingston and Speights with Thompson, Barnes, and Iguodala (70 possessions). Finally, it’s also interesting to note that Livingston appears to pair better with other reserves than with starters (save Draymond Green, further highlighting Green’s overall value), an observation that raises important questions about how players understand and occupy particular roles on the team.

Table 9 shows the best and worst triples with at least 200 possessions. The grouping of Green–Curry–Thompson is far and away the most dominant triple, and safely (and unsurprisingly) earns designation as the Warriors’ big three. Other notable triples include starters like Green and Curry or Green and Thompson together with Andre Iguodala who came off the bench, and more lightly used triples like Curry–Barbosa–Speights who had an SCLP of 4.6 in 245 possessions. Analyzing subpairs of these groups shows a better stacking of synergies in the triples that include Iguodala—he pairs well with Green, Curry, and Thompson in the second order space as well, while either of Barbosa or Speights paired poorly with Curry. Still, Barbosa with Speights was quite strong as a pair, and we see that the addition of Curry does provide added value as a pure triple. Interesting ineffective triples include Iguodala and Bogut with either of Curry or Green, especially in light of the fact that Bogut–Iguodala was

also a weak pairing (see detailed tables in the appendix).

Figure 1 shows that the most effective player-triples as identified by spectral analysis are positively correlated with higher values of plus-minus. As raw group plus-minus decreases,

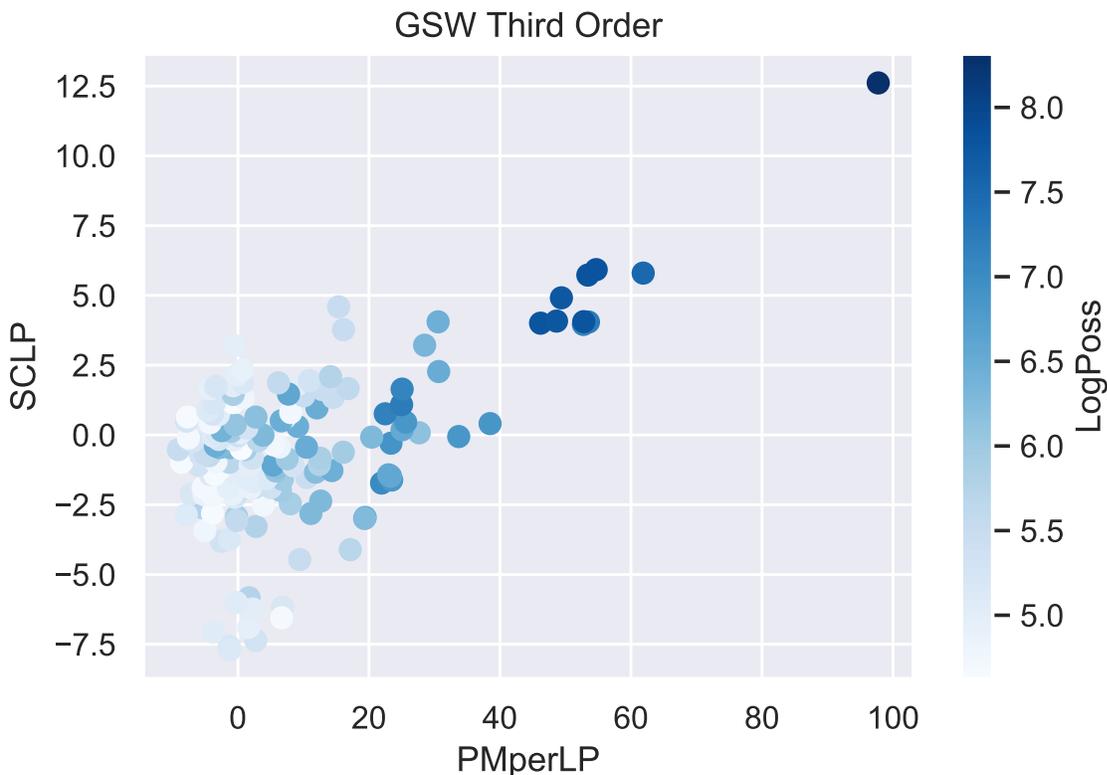


Figure 1: Third-order effects for triples with more than 100 possessions the 2015-2016 Golden State Warriors. The x -axis gives the group's plus-minus per log possession (PMperLP) while the y -axis shows the spectral contribution per log possession (SCLP). Observations are shaded by number of possessions.

however, we see considerable variation in the spectral contributions of the groups (and in number of possessions played). This suggests the following narrative: while it may be relatively easy to identify the team's top groups, it is considerably more difficult to identify positive and negative synergies among the remaining groups, especially when controlling for lower-order contributions. Spectral analysis suggests several opportunities for constructing more optimal lineups with potential for untapped competitive advantage, especially when more obvious dominant groupings are unavailable.

Table 10 shows top and bottom three third-order effects for the 15-16 Boston Celtics. (The appendix includes more complete tables for Boston including effects of all orders.) Figure 2 gives contrasting bar plots of the third-order effects for both Boston and Golden State. The Celtics have fewer highly dominant groups. In particular, we note that the

P1	P2	P3	SCLP	PM	Poss
Evan Turner	Kelly Olynyk	Jonas Jerebko	2.9	110.1	879
Isaiah Thomas	Avery Bradley	Jared Sullinger	2.7	177.7	2642
Avery Bradley	Jae Crowder	Jared Sullinger	2.3	139.3	2216
Isaiah Thomas	Evan Turner	Kelly Olynyk	-1.8	-30.9	870
Avery Bradley	Jared Sullinger	Jonas Jerebko	-2.3	-11.7	194
Isaiah Thomas	Avery Bradley	Jonas Jerebko	-2.4	-1.6	290

Table 10: Top and bottom three third-order effects for BOS with at least 150 possessions.

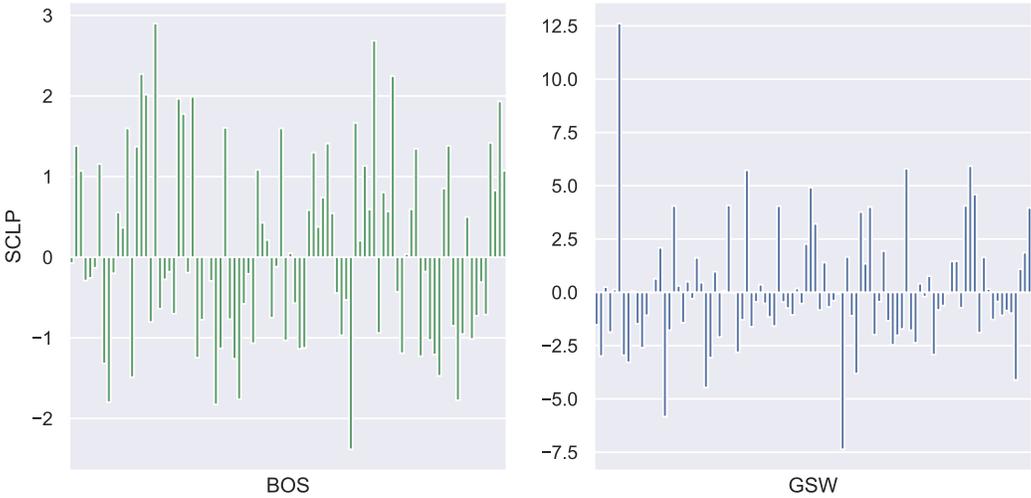


Figure 2: Bar graph of third order spectral contributions per log possession (SCLP) for BOS and GSW for groups with more than 150 possessions.

spectral signature of the Celtics is distinctly different from that of the Warriors in that Boston lacks anything resembling the big-three of Golden State. While SCLP values are not directly comparable across teams (they depend, for instance, on the norm of the overall

team success function when projected into each effect space), the relative values within an effect-space are comparable. Similarly, the SCLP values also depend on the norm of the interpretable vector used in Mallows’s method. As a result, the values are not directly comparable across effect spaces— a problem we return to below.

In fourth and fifth-order spaces the numbers of high-possession groups begins to decline, as alluded to above. (See appendix for complete tables.) Still, it is interesting to note that spectral analysis flags the Warriors small lineup of Green–Curry–Thompson–Barnes–Iguodala as the team’s best, even over the starting lineup with Bogut replacing Barnes. It also prefers two lesser-used lineups to the Warriors’ second most-used lineup of Green–Curry–Thompson–Bogut–Rush. Also of note is the fact that Golden State’s best group of three and best group of four are both subsets of the starting lineup— another instance of stacking of positive effects—while neither of Boston’s best groups of three or four are part of their starting lineup.

6 Connection With Linear Models

Before moving on, we consider the connection between spectral analysis and a related approach via linear regression which will likely be more familiar to the sports analytics community.

Recalling our assumption of a 15 man roster, consider the problem of modeling a lineup’s plus-minus, given by $f(L)$ for lineup L , using indicator variables that correspond to all possible groups of players. Label the predictor variables X_1, X_2, \dots, X_p , where each variable corresponds to a group of players (with some fixed group order). Thus, the variable X_i is 1 when the players from group i are on the floor, and zero otherwise. If the first fifteen variables are the indicator functions of the individual players X_1, X_2, \dots, X_{15} , then the group variables, the X_i for $i > 15$, are interaction terms. For instance, the variable corresponding to the group $\{1, 2, 3\}$ is $X_1X_2X_3$. This approach is therefore similar to an adjusted plus minus with interactions approach. Including all possible group effects, however, means that the number of predictors is quite large and depending on the number of observations, we may be in a situation where $p \gg N$. Moreover, the nature of player usage in lineups means that there is a significant multicollinearity issue. Consequently, an attempt to quantify group effects in a regression model of this sort will rely on a shrinkage technique like ridge regression.

Let N be the number of lineups, and $y = f(L)$, an $N \times 1$ column vector. Let \mathbf{X} be the $N \times (p + 1)$ matrix whose first column is the vector of all ones and where the i -th row consists of the binary value of each predictor variable for the i -th player group. The vector of ridge coefficients $\hat{\beta}^{\text{ridge}}$ minimizes the penalized residual sum of squares:

$\arg \min_{\beta} \{ \|y - \mathbf{X}\beta\|^2 + \lambda \sum_{i=1}^p \beta_i^2 \}$. The non-negative parameter λ serves as a penalty on the L_2 -norm of the solution vector. (The intercept is not included in the ridge penalty.) The ridge approach reduces the variability exhibited by the least squares coefficients in the presence of multicollinearity by shrinking the coefficient estimates in the model towards zero (and toward each other). One can show that ridge regression uses the singular values of the covariance matrix associated with the centered version of \mathbf{X} to disproportionately shrink coefficients associated with inputs where the data exhibits lower degrees of variance. See Friedman et al. (2001) for details.

The fitted coefficients $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$ in the ridge regression model attempt to measure the contribution of group i while controlling for the contributions of all other groups and individuals. We note that this modeling approach resembles work in Sill (2010), Grasseti et al. (2019a), and Grasseti et al. (2019b), though there are key differences which we explore below. In particular, note that we model group contributions aggregated over all opponents, and without controlling for the quality of the opponents faced. This simplified approach allows for a more direct comparison with the results of spectral analysis above.

Tables 11 and 12 give the ridge regression coefficients associated with the top 5 individuals, pairs, and triples for the Warriors. Comparing with Tables 5, 6, and 9 shows both

Individual	Estimate	P1	P2	Pair Estimate
Draymond Green	0.28	Draymond Green	Stephen Curry	0.65
Stephen Curry	0.25	Stephen Curry	Andrew Bogut	0.53
Klay Thompson	0.15	Stephen Curry	Klay Thompson	0.47
Andrew Bogut	0.14	Draymond Green	Klay Thompson	0.47
Festus Ezeli	0.02	Draymond Green	Andrew Bogut	0.46

Table 11: Best individuals and pairs using the linear model.

P1	P2	P3	Estimate
Draymond Green	Stephen Curry	Andrew Bogut	1.61
Stephen Curry	Klay Thompson	Andrew Bogut	1.49
Draymond Green	Stephen Curry	Klay Thompson	1.39
Draymond Green	Klay Thompson	Andrew Bogut	1.24
Draymond Green	Klay Thompson	Harrison Barnes	1.03

Table 12: Top triples according to the linear model.

some overlap in the top rated groups, but also significant differences with respect to both ordering and magnitude of contribution. In particular, the linear model appears to value the

contributions of Andrew Bogut considerably more than spectral analysis. It is also notable that spectral analysis identifies a clearly dominant big three of Green–Curry–Thompson, in contrast to the considerably different result arising from the modeling approach which ranks that group third.

We can interpret the linear model determined by $\hat{\beta}^{\text{ridge}}$ as giving a similar decomposition to the spectral decomposition in (2). For each lineup L we have predicted success given by

$$\hat{y} = \mathbf{X}_L \hat{\beta}^{\text{ridge}} \quad (8)$$

where \mathbf{X}_L is now the $\binom{15}{5} \times (p+1)$ matrix whose first column is all 1s, and whose $i, j+1$ entry is 1 if the j -th player group is part of the i -th lineup. (We have fixed a particular ordering of lineups.) The columns of \mathbf{X}_L (the X_i) that correspond to individual players can be understood as spanning a subspace W_1 analogous to V_1 in (2). Similarly, W_2 is spanned by the columns of \mathbf{X}_L corresponding to pair interactions, and so on for all groups through full five player lineups. The particular linear combinations in each W_i determined by the respective coordinates of $\hat{\beta}^{\text{ridge}}$ are analogous to the $\mathbf{pr}_{V_i} f$. In fact, the space of all lineup functions can be written

$$V = W_0 + W_1 + W_2 + W_3 + W_4 + W_5, \quad (9)$$

where W_i is the space of interaction effects for groups of size i .

Still, there are important differences between (2) and (9). While V_0 and W_0 are both one-dimensional, for $i \geq 1$ the dimensions of the W_i are strictly larger than those of their V_i counterparts. For instance, W_5 includes a vector for each possible set of five players from the original fifteen. Similarly W_4 and groups of four, and so on. Thus, the dimension of W_5 is 3003 (the number of lineups), which is the same as the dimension of V itself. By contrast the dimension of V_5 in (2) is only 1638. Similarly the dimension of W_4 is 1365 while that of V_4 is 350. Clearly, the decomposition in (8) is highly non-orthogonal (explaining the $+$ rather than \oplus notation). It is easy to find vectors in W_i that overlap with W_j in the sense that their inner product is non-zero. In the context of basketball, the contribution of a group of, for example, 5 players is not necessarily separate from a constituent group of four (or any other number of) players despite the use of shrinkage methods.

The decomposition in (2) is special in that it gives minimal subspaces that are invariant under relabeling and mutually orthogonal as described in section 3. As we've seen, spectral analysis achieves this at the expense of easy interpretation of group contributions. This is a drawback to spectral analysis that (8) does not have, and is an appealing feature of regression models. The interaction term associated with a group of i players in a regression model is easy to understand. Still, as we see above one must balance either ease of interpretation, or orthogonality of effects.

7 Stability

In this section we take a first step to addressing questions of the stability of spectral analysis. We seek evidence that spectral analysis is indicative of a true signal, and that should the data have turned out slightly differently, the analysis would not change dramatically. Since spectral analysis works on the lineup function $f(L)$, which is aggregated over all of a team's plays involving L , we need to introduce variability into the values of $f(L)$. A fully aggregated NBA season is, in a sense, a complete record of all events and lineup outcomes in that season. Still, it seems reasonable to leverage the variability inherent in the many observed results of a lineup's plays, as well as the substitution patterns of coaches, and suggest a bootstrapping approach.

To that end, we start with the actual 15-16 season for the Boston Celtics. We can then build a bootstrapped season by sampling plays, with replacement, from the set of all plays in the actual season. (We sample the same number of plays as in the actual season.) A play is defined as a connected sequence of events surrounding a possession in the team's play-by-play data. For example, a play might involve a sequence like a missed shot, offensive rebound, and a made jump shot; or, a defensive rebound followed by a bad pass turnover. When sampling from a team's plays, a particular lineup will be selected with a probability proportional to the number of plays in which that lineup participated. We generate 500 bootstrapped seasons, process each using the methodology of sections 2 and 3 to produce success functions f_{boot} , and then apply spectral analysis to each. We thus have a bootstrapped distribution of lineup plus-minus and possession values over each lineup L , which in turn gives plus-minus and possession distributions of all player-groups. While the the number of possessions played is highly stable for both full-lineups and smaller player-groups, there is considerable variability in plus-minus values over the bootstrapped seasons. Lineups with a significant number of possessions exhibit both positive and negative performance, and the balance between the positive and negative plays is delicate.

The variability in group PM presents a challenge in gauging the stability of the spectral analysis associated with a player group. Take, for example, the Thomas–Bradley–Crowder triple for the Celtics. The actual season's plus-minus for this group was 154.8 in 2572 possessions. Over the bootstrapped seasons the group has means of 145.9 and 2574.1 for plus-minus and possessions, respectively. On the other hand, the standard deviation of the plus-minus values is 82.8 versus only 47.7 for possessions. Thus, some of the variability in the spectral contribution of the group over the bootstrapped seasons should be expected since, in fact, the group was less effective in some of those seasons. Figure 3 shows SCLP plotted against PMperLP for the Thomas–Bradley–Crowder triple in 500 bootstrapped seasons. Of course, spectral analysis purports to do more than raw plus-minus by removing otherwise confounding colinearities and overlapping effects. Not surprisingly, therefore, we still see

variability in SCLP within a band of plus-minus values, but the overall positive correlation, whereby SCLP increases in seasons where the group tended to outscore its opponents, is reasonable.

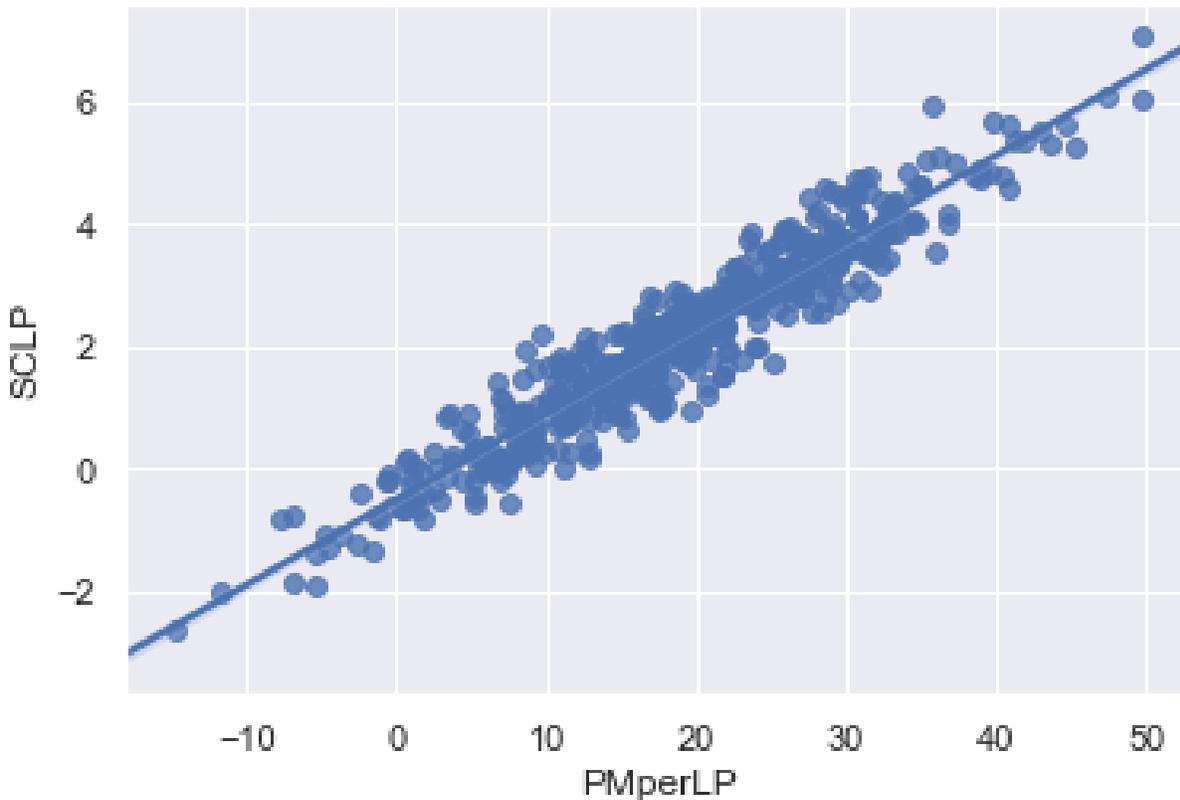


Figure 3: Spectral contribution per log possession (SCLP) versus plus-minus per log possession (PMperLP) for Thomas–Bradley–Crowder triple in 500 bootstrapped seasons. Each bootstrapped season consists of sampling plays (connected sequences of game events) with replacement from the set of all season plays. Resampled season data is then processed as in section 2 and group contributions are computed via spectral analysis as in section 3.

Also intuitively, the strength of the correlation between group plus-minus and spectral contribution depends on the number of possessions played. Fewer possessions means that group’s contribution is more dependent on other groups and hence exhibits more variability. The mean possessions for the Thomas–Bradley–Crowder triple in Fig.3 is 2574, and has a Pearson correlation of $r = 0.953$. The group Thomas–Turner–Zeller, on the other hand,

has $r = 0.688$ with a mean of 305 possessions. A group like Jared Sullinger–Marcus Smart is particularly interesting. This pair has a season plus-minus of 25.0 in 1116 possessions. In 500 bootstrap seasons, they have a mean plus-minus of 23.6 and mean possessions of 1118.3. The value of the group’s plus-minus is negative in only 32.4% of those seasons. Should this group, therefore, be considered effective overall? Spectral analysis answers with a fairly emphatic *no*. After removing other group contributions their SCLP as a pure pair is negative in 90.6% of bootstrapped seasons, while still exhibiting strong correlation with overall plus-minus ($r = 0.73$). Similarly, the Bradley–Smart pair has a season plus-minus of 45.3 in 1679 possessions. In 500 bootstrap seasons, they have a mean plus-minus of 40.4 and mean possessions of 1679. Their plus-minus is negative in 27% of those seasons while their spectral contribution is negative in 81% of bootstrapped seasons.

8 Importance of Effect Spaces

Another natural question is how to value the relative importance of the group-effect spaces. One way to gauge importance uses the squared L_2 norm of the success function in each space. Since the spaces are mutually orthogonal, we have $\|f\|^2 = \|f_1\|^2 + \|f_2\|^2 + \|f_3\|^2 + \|f_4\|^2 + \|f_5\|^2$. (Recall that f_i is the projection of f onto the i -th order effect space V_i .) One can then measure the total mass of f that is concentrated in each effect space. For example, if we found that the mass of the success function was concentrated in the mean space, and thus, a constant function gave a good approximation to f , we could conclude that the particular lineup used by this team was largely irrelevant—the success of the team never strayed far from the mean and was not strongly affected by any groups. This would be an easy team to coach. Of course, this is not the case in basketball, as evidenced by the L_2 norm squared distribution of the sample of teams in Table 13.

By this measure, the higher-order spaces are dominant as they hold most of the mass of the success function. An issue with this metric, however, is the disparity in the dimensions of the spaces. Because V_5 is 1638-dimensional, we might expect the mass of f to be disproportionately concentrated in that space. In fact, a random unit vector projected into each of the effect spaces would be, on average, distributed according to the null distribution in Table 13, with mass proportional to the dimension of each of the spaces in question.

Moreover, we can take the true success function of a team and break the dependence on the actual player groups as follows. Recall that the raw data f records the plus-minus for each of the possible 3003 lineups. We then take f and randomly permute the values so that there is no connection between the lineup and the value associated with that lineup. Still, however, the overall plus-minus and mean of f are preserved. We can then run spectral analysis on the permuted f and record the distribution of the squared L_2 norm in each space.

Team	V_0	V_1	V_2	V_3	V_4	V_5
BOS	0.001	0.012	0.048	0.138	0.297	0.504
CLE	0.003	0.021	0.058	0.150	0.301	0.467
GSW	0.003	0.031	0.092	0.203	0.312	0.360
HOU	0.000	0.007	0.037	0.123	0.285	0.548
OKC	0.001	0.011	0.038	0.137	0.304	0.510
POR	0.000	0.004	0.027	0.112	0.289	0.568
SAS	0.007	0.027	0.072	0.173	0.294	0.427
Null	0.000	0.005	0.03	0.117	0.303	0.545

Table 13: Distribution of the squared L_2 -norm of the team success function over the effect spaces.

Space	BOS	GSW
First	0.005	0.005
Second	0.030	0.030
Third	0.117	0.116
Fourth	0.302	0.302
Fifth	0.543	0.544

Table 14: Average fraction of squared L_2 mass by order effect space using randomly permuted success function.

Repeating this experiment 500 times for both GSW and BOS give means in Table 14 that exactly conform to the null distribution in Table 13.

An alternative measure of the importance of each effect space is given by measuring the extent to which projections onto V_i deviate from the null distribution. By this measure of importance, there is some preliminary evidence that strong teams shift the mass of f from V_5 into lower-order spaces, particularly V_1 , V_2 , and V_3 . This is interesting as it agrees with the idea that building an elite team requires a group of three stars. Using all 30 NBA teams, we compute correlations of $r = 0.51$, $r = 0.58$ and $r = 0.55$, respectively, between win-percentage and the projected mass f in the first, second, and third-order spaces. Win-percentage and fifth-order projection have correlation coefficient $r = -0.54$. As pointed out in Diaconis (1989), however, care must be taken when looking at deviation from the null distribution if the projections are highly structured and lie close to a few of the interpretable vectors. This is a direction for further inquiry.

9 Conclusion

Spectral analysis proposes a new approach to understanding and quantifying group effects in basketball. By thinking of the success of a team as function on lineups, we can exploit the structure of functions on permutations to decompose the team success function. The resulting Fourier expansion is naturally interpreted as quantifying the group effects to overall team success. The resulting analysis brings insight into important and difficult questions like which groups of players work effectively together, and which do not. Furthermore, the spectral analysis approach is unique in addressing questions of lineup synergies by presenting an EDA summary of the actual team data without making the kind of modeling or skill-based assumptions of other methods.

There are several directions for future work. First, the analysis presented used raw lineup level plus-minus to measure success. This approach has the advantage of keeping the analysis tethered to data that is intuitive, and helps avoid pitfalls arising from low-possession lineups. Still, adjusting the lineup level plus-minus to account for quality of opponent, for example, seems like a valuable next step. Another straight forward adjustment to raw plus-minus data would involve devaluing so-called garbage time possessions when the outcome of the game is not in question.

As presented here, spectral analysis provides an in-depth exploratory analysis of a team's lineups. Still, the results of spectral analysis could also add valuable inputs to more traditional predictive models or machine learning approaches to projecting group effects. Similarly, it would be interesting to use spectral analysis as a practical tool for lineup suggestions. While the orthogonality of the spectral decomposition facilitates valuation of pure player-groups, the question of lineup construction realistically begins at the level of individuals and works up, hopefully stacking the contributions of individuals with strong pairs, triples, and so-on. A strong group of three, for instance, without any strong individual players may be interesting from an internal development perspective, or at the edges of personnel utility, but may also be of limited practical value from the perspective of constructing a strong lineup. Development of a practical tool would likely require further analysis of the ideas in sections 7 and 8 based on ideas in Diaconis et al. (1998). For example, given data (a function on lineups), we might fix the projection of that data onto certain spaces (like the first or second order), and then generate new sample data conditional on that fixed projection. The resulting projections in the higher-order spaces would give some evidence for how the fixed lower-order projections affect the mass of f in the higher-order effects spaces. This would help give a more detailed sense of variability of projections, and a more definitive answer to the question of which spaces are most important, and how the spectral signature of a team correlates with team success. With that information in place, however, one can build tools to suggest lineup replacements that maximize the stacking of a team's most important

groups.

References

- Basketball-Reference. Glossary. <https://www.basketball-reference.com/about/glossary.html>. "[Online; accessed 17-May-2019]".
- Persi Diaconis. Group representations in probability and statistics. *Lecture Notes-Monograph Series*, 11:pp. i–vi+1–192, 1988. ISSN 07492170. URL <http://www.jstor.org/stable/4355560>.
- Persi Diaconis. A generalization of spectral analysis with application to ranked data. *The Annals of Statistics*, 17(3):pp. 949–979, 1989. ISSN 00905364. URL <http://www.jstor.org/stable/2241705>.
- Persi Diaconis, Bernd Sturmfels, et al. Algebraic algorithms for sampling from conditional distributions. *The Annals of statistics*, 26(1):363–397, 1998.
- David Steven Dummit and Richard M. Foote. *Abstract algebra*. John Wiley & sons, Hoboken, NJ, 2004. ISBN 0-471-43334-9. URL <http://opac.inria.fr/record=b1133479>.
- Jerome Friedman, Trevor Hastie, and Robert Tibshirani. *The elements of statistical learning*, volume 1. Springer series in statistics Springer, Berlin, 2001.
- Luca Grassetti, Ruggero Bellio, Giovanni Fonseca, and Paolo Vidoni. Estimation of lineup efficiency effects in basketball using play-by-play data. In G. Arbia, S. Peluso, A. Pini, and G. Rivellini, editors, *Book of Short Papers SIS2019*. Pearson, 2019a.
- Luca Grassetti, Ruggero Bellio, Giovanni Fonseca, and Paolo Vidoni. Play-by-play data analysis for team managing in basketball. In Dimitris Karlis, Ioannis Ntzoufras, and Sotiris Drikos, editors, *Proceedings of MathSport International 2019 Conference (e-book)*. Propobos Publications, 2019b.
- Giuseppe Jurman, Stefano Merler, Annalisa Barla, Silvano Paoli, Antonio Galea, and Cesare Furlanello. Algebraic stability indicators for ranked lists in molecular profiling. *Bioinformatics*, 24(2):258–264, 2008. doi: 10.1093/bioinformatics/btm550. URL <http://bioinformatics.oxfordjournals.org/content/24/2/258.abstract>.
- R. Kakarala. A signal processing approach to fourier analysis of ranking data: The importance of phase. *Signal Processing, IEEE Transactions on*, 59(4):1518–1527, April 2011. ISSN 1053-587X. doi: 10.1109/TSP.2010.2104145.

- Risi Kondor and Walter Dempsey. Multiresolution analysis on the symmetric group. In F. Pereira, C.J.C. Burges, L. Bottou, and K.Q. Weinberger, editors, *Advances in Neural Information Processing Systems 25*, pages 1637–1645. Curran Associates, Inc., 2012. URL <http://papers.nips.cc/paper/4720-multiresolution-analysis-on-the-symmetric-group.pdf>.
- Risi Kondor, Andrew Howard, and Tony Jebar. Multi-object tracking with representations of the symmetric group. 2:211–218, 2007.
- Joseph Kuehn. Accounting for complementary skill sets when evaluating nba players? values to a specific team. In *2016 MIT Sloan Sports Analytics Conference*, volume 6, 2016.
- Brian L. Lawson, Michael E. Orrison, and David T. Uminsky. Spectral analysis of the supreme court. *Mathematics Magazine*, 79(5):pp. 340–346, 2006. ISSN 0025570X. URL <http://www.jstor.org/stable/27642969>.
- David K Maslen, Michael E Orrison, and Daniel N Rockmore. Computing isotypic projections with the lanzos iteration. *SIAM Journal on Matrix Analysis and Applications*, 25(3):784–803, 2003.
- Allan Z Maymin, Philip Z Maymin, and Eugene Shen. Nba chemistry: Positive and negative synergies in basketball. *International Journal of Computer Science in Sport*, 12(2):4–23, 2013.
- Krishna P. Paudel, Mahesh Pandit, and Michael A. Dunn. Using spectral analysis and multinomial logit regression to explain households’ choice patterns. *Empirical Economics*, 44(2):739–760, 2013. ISSN 0377-7332. doi: 10.1007/s00181-012-0558-4. URL <http://dx.doi.org/10.1007/s00181-012-0558-4>.
- Jean-Pierre Serre. *Linear representations of finite groups*, volume 42. Springer Science & Business Media, 2012.
- J. Sill. Improved nba adjusted +/- using regularization and out-of-sample testing. In *MIT Sloan Sports Analytics Conference*. 2010.
- D. Uminsky, M. Banuelos, L. Gonzalez-Albino, R. Garza, and S. A. Nwakanma. Detecting higher order genomic variant interactions with spectral analysis. In *2019 27th European Signal Processing Conference (EUSIPCO)*, pages 1–5, Sep. 2019. doi: 10.23919/EUSIPCO.2019.8902725.

David Uminsky, Rosa Garza, Lillian González, Sylvia Akueze Nwakanma, Stephen Devlin, and Mario Banuelos. Detecting higher order variable interactions: A spectral analysis approach. In *LatinX in AI Workshop at NeurIPS*. 2018.

10 Appendix

We include more detailed tables reporting group effects of all orders for both Golden State and Boston.

Player	SCLP	PM	Poss
Draymond Green	17.2	1038.4	5800
Stephen Curry	15.9	978.7	5610
Klay Thompson	12.0	808.6	5453
Andre Iguodala	3.5	436.1	3516
Andrew Bogut	2.8	403.6	2951
Harrison Barnes	2.2	384.7	4138
Festus Ezeli	-1.9	225.0	1550
Shaun Livingston	-2.0	211.1	2980
Leandro Barbosa	-5.8	70.6	2144
Brandon Rush	-7.1	23.3	2087
Marreese Speights	-7.4	20.0	1630
Ian Clark	-9.8	-51.9	1108
Anderson Varejao	-11.1	-34.4	368
Jason Thompson	-11.2	-33.8	339
James Michael McAdoo	-12.1	-85.0	526

Table 15: All first order effects for GSW.

P1	P2	SCLP	PM	Poss
Draymond Green	Stephen Curry	13.3	979.9	5102
Stephen Curry	Klay Thompson	11.2	827.8	4311
Draymond Green	Klay Thompson	11.1	847.8	4678
Leandro Barbosa	Marreese Speights	5.3	76.2	983
Draymond Green	Andre Iguodala	4.3	490.0	2165
Klay Thompson	Andre Iguodala	4.2	411.4	1764
Stephen Curry	Andre Iguodala	3.9	460.0	2185
Klay Thompson	Harrison Barnes	3.9	396.0	3058
Klay Thompson	Andrew Bogut	3.7	394.3	2637
Leandro Barbosa	Ian Clark	3.5	-9.6	325
Draymond Green	Harrison Barnes	3.3	445.1	2634
Stephen Curry	Harrison Barnes	3.2	423.3	2809
Marreese Speights	Ian Clark	2.5	-44.2	493
Harrison Barnes	Andrew Bogut	2.1	206.2	1527
Leandro Barbosa	Brandon Rush	1.8	-22.4	638
Brandon Rush	Ian Clark	1.7	-64.6	463.0
Andre Iguodala	Festus Ezeli	1.7	152.6	999.0
Stephen Curry	Andrew Bogut	1.6	378.5	2530.0
Shaun Livingston	Marreese Speights	1.6	17.8	1014.0
Leandro Barbosa	Festus Ezeli	1.3	26.2	468.0
Stephen Curry	Brandon Rush	-2.5	140.7	1260.0
Harrison Barnes	Marreese Speights	-2.9	-48.8	794.0
Draymond Green	Brandon Rush	-3.0	144.4	1266.0
Klay Thompson	Festus Ezeli	-3.1	138.9	824.0
Harrison Barnes	Brandon Rush	-3.1	-50.1	546.0
Harrison Barnes	Festus Ezeli	-3.2	9.5	598
Draymond Green	Leandro Barbosa	-3.3	154.7	860
Klay Thompson	Shaun Livingston	-3.6	111.8	1412
Stephen Curry	Leandro Barbosa	-3.7	126.3	883
Draymond Green	Marreese Speights	-4.0	129.7	492
Andre Iguodala	Brandon Rush	-4.4	-59.9	399
Klay Thompson	Ian Clark	-5.1	10.3	498
Stephen Curry	Marreese Speights	-5.7	73.5	423
Klay Thompson	Marreese Speights	-6.4	-5.1	581
Klay Thompson	James Michael McAdoo	-7.1	-28.9	241
Draymond Green	Ian Clark	-7.2	33.3	424
Klay Thompson	Leandro Barbosa	-7.2	4.8	349
Stephen Curry	Ian Clark	-8.1	14.0	220
Draymond Green	Anderson Varejao	-9.5	7.2	217
Stephen Curry	Anderson Varejao	-10.1	-26.9	237

Table 16: Second order effects for GSW with at least 200 possessions.

P1	P2	P3	SCLP	PM	Poss
Draymond Green	Stephen Curry	Klay Thompson	12.6	812.7	4085
Draymond Green	Klay Thompson	Harrison Barnes	5.9	427.3	2473
Draymond Green	Stephen Curry	Andre Iguodala	5.8	464.8	1830
Stephen Curry	Klay Thompson	Harrison Barnes	5.7	416.5	2431
Stephen Curry	Klay Thompson	Andrew Bogut	4.9	382.2	2296
Stephen Curry	Leandro Barbosa	Marreese Speights	4.6	84.6	245
Draymond Green	Stephen Curry	Andrew Bogut	4.1	377.4	2346
Draymond Green	Stephen Curry	Harrison Barnes	4.1	411.3	2421
Stephen Curry	Andre Iguodala	Festus Ezeli	4.1	197.2	633
Draymond Green	Klay Thompson	Andre Iguodala	4.0	388.4	1418
Draymond Green	Klay Thompson	Andrew Bogut	4.0	359.8	2409
Stephen Curry	Klay Thompson	Andre Iguodala	4.0	377.0	1270
Draymond Green	Leandro Barbosa	Marreese Speights	3.8	88.9	248
Draymond Green	Andre Iguodala	Festus Ezeli	3.2	180.8	569
Klay Thompson	Harrison Barnes	Andre Iguodala	2.3	199.5	671
Leandro Barbosa	Marreese Speights	Ian Clark	-2.1	-0.2	203
Draymond Green	Andre Iguodala	Andrew Bogut	-2.4	79.4	535
Draymond Green	Shaun Livingston	Andrew Bogut	-2.5	47.4	370
Klay Thompson	Harrison Barnes	Marreese Speights	-2.6	-31.4	323
Stephen Curry	Andre Iguodala	Andrew Bogut	-2.8	70.2	541
Klay Thompson	Harrison Barnes	Festus Ezeli	-2.9	-1.2	353
Draymond Green	Stephen Curry	Leandro Barbosa	-3.0	126.9	687
Stephen Curry	Klay Thompson	Shaun Livingston	-3.0	121.3	530
Klay Thompson	Harrison Barnes	Brandon Rush	-3.1	-1.0	265
Draymond Green	Harrison Barnes	Festus Ezeli	-3.3	16.0	326
Stephen Curry	Andre Iguodala	Brandon Rush	-3.8	-13.5	207
Draymond Green	Stephen Curry	Marreese Speights	-4.1	97.9	299
Draymond Green	Klay Thompson	Marreese Speights	-4.5	52.2	250
Draymond Green	Klay Thompson	Ian Clark	-5.8	9.8	316
Draymond Green	Stephen Curry	Ian Clark	-7.4	14.5	205

Table 17: Third order effects for GSW with at least 200 possessions.

P1	P2	P3	P4	SCPLP	PM	Poss
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	8.7	401.6	2271
Draymond Green	Stephen Curry	Klay Thompson	Andrew Bogut	7.8	365.7	2159
Draymond Green	Stephen Curry	Klay Thompson	Andre Iguodala	7.7	364.9	1157
Draymond Green	Andre Iguodala	Shaun Livingston	Festus Ezeli	3.9	76.8	201
Draymond Green	Stephen Curry	Leandro Barbosa	Marreese Speights	3.9	67.8	173
Draymond Green	Stephen Curry	Andre Iguodala	Festus Ezeli	3.8	170.3	526
Stephen Curry	Klay Thompson	Harrison Barnes	Andre Iguodala	3.3	171.3	451
Draymond Green	Stephen Curry	Harrison Barnes	Andrew Bogut	2.8	162.3	1165
Stephen Curry	Andre Iguodala	Shaun Livingston	Festus Ezeli	2.7	64.9	201
Draymond Green	Stephen Curry	Klay Thompson	Brandon Rush	2.4	177.9	870
Draymond Green	Stephen Curry	Harrison Barnes	Andre Iguodala	2.3	157.7	419
Draymond Green	Klay Thompson	Harrison Barnes	Andre Iguodala	2.0	158.0	417
Draymond Green	Klay Thompson	Harrison Barnes	Andrew Bogut	1.8	158.4	1221
Draymond Green	Stephen Curry	Shaun Livingston	Festus Ezeli	1.7	75.7	198
Stephen Curry	Klay Thompson	Harrison Barnes	Andrew Bogut	1.5	154.0	1235
Draymond Green	Stephen Curry	Andre Iguodala	Andrew Bogut	-1.3	79.3	433
Draymond Green	Stephen Curry	Harrison Barnes	Shaun Livingston	-1.7	57.4	214
Harrison Barnes	Andre Iguodala	Shaun Livingston	Festus Ezeli	-2.0	-20.7	160
Draymond Green	Stephen Curry	Klay Thompson	Shaun Livingston	-2.1	116.5	485
Draymond Green	Stephen Curry	Harrison Barnes	Brandon Rush	-2.1	23.0	160
Draymond Green	Klay Thompson	Harrison Barnes	Festus Ezeli	-2.3	17.1	299
Stephen Curry	Klay Thompson	Harrison Barnes	Brandon Rush	-2.4	30.3	152
Stephen Curry	Klay Thompson	Harrison Barnes	Festus Ezeli	-2.5	19.0	309
Draymond Green	Stephen Curry	Harrison Barnes	Festus Ezeli	-3.0	17.7	309
Draymond Green	Klay Thompson	Andre Iguodala	Shaun Livingston	-3.0	18.2	261

Table 18: Fourth order effects for GSW with at least 150 possessions.

P1	P2	P3	P4	P5	SCLP	PM	Poss
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	Andre Iguodala	10.3	152.9	372
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	Andrew Bogut	7.0	142.0	1140
Draymond Green	Stephen Curry	Andre Iguodala	Shaun Livingston	Festus Ezeli	6.9	67.3	160
Draymond Green	Stephen Curry	Klay Thompson	Andrew Bogut	Brandon Rush	5.2	88.2	535
Draymond Green	Stephen Curry	Klay Thompson	Andre Iguodala	Andrew Bogut	5.1	98.2	310
Draymond Green	Stephen Curry	Klay Thompson	Andre Iguodala	Festus Ezeli	4.9	85.7	266
Draymond Green	Stephen Curry	Klay Thompson	Shaun Livingston	Andrew Bogut	1.9	42.7	112
Harrison Barnes	Shaun Livingston	Leandro Barbosa	Brandon Rush	Marreese Speights	1.8	6.4	87
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	Shaun Livingston	0.7	42.7	175
Harrison Barnes	Andre Iguodala	Shaun Livingston	Leandro Barbosa	Marreese Speights	-0.8	-3.1	172
Harrison Barnes	Andre Iguodala	Shaun Livingston	Leandro Barbosa	Festus Ezeli	-1.3	-9.9	102
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	Festus Ezeli	-1.9	20.1	283
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	Brandon Rush	-2.3	28.0	123
Draymond Green	Stephen Curry	Klay Thompson	Andre Iguodala	Shaun Livingston	-5.1	2.3	98
Draymond Green	Stephen Curry	Klay Thompson	Harrison Barnes	James Michael McAdoo	-5.9	-14.2	91

Table 19: Fifth order effects for GSW with at least 80 possessions.

P1	P2	SCLP	PM	Poss
Andrew Bogut	Klay Thompson	3.7	394.3	2637
Andrew Bogut	Harrison Barnes	2.1	206.2	1527
Andrew Bogut	Stephen Curry	1.6	378.5	2530
Andrew Bogut	Draymond Green	0.8	371.5	2596
Andrew Bogut	Brandon Rush	0.7	54.5	733
Andrew Bogut	Ian Clark	-0.3	6.1	198
Andrew Bogut	Shaun Livingston	-0.6	77.4	573
Andrew Bogut	Leandro Barbosa	-1.6	16.3	166
Andrew Bogut	Andre Iguodala	-2.1	107.0	785

Table 20: Pairs involving Andrew Bogut (with at least 150 possessions).

P1	P2	SCLP	PM	Poss
Shaun Livingston	Anderson Varejao	2.0	-1.5	174
Shaun Livingston	Marreese Speights	1.6	17.8	1014
Shaun Livingston	Draymond Green	1.2	323.6	1486
Shaun Livingston	Ian Clark	0.9	-25.7	378
Shaun Livingston	Leandro Barbosa	0.9	15.2	1210
Shaun Livingston	James Michael McAdoo	0.8	-41.6	180
Shaun Livingston	Festus Ezeli	0.4	49.0	654
Shaun Livingston	Stephen Curry	-0.1	265.5	1120
Shaun Livingston	Andrew Bogut	-0.6	77.4	573
Shaun Livingston	Harrison Barnes	-1.1	55.2	1475
Shaun Livingston	Andre Iguodala	-1.3	65.2	1605
Shaun Livingston	Brandon Rush	-1.5	-63.2	536
Shaun Livingston	Klay Thompson	-3.6	111.8	1412

Table 21: Pairs involving Shaun Livingston (with at least 150 possessions).

Player 1	Player 2	Player 3	SCLP	PM	Poss
Klay Thompson	Harrison Barnes	Shaun Livingston	-1.1	35.3	733
Draymond Green	Andre Iguodala	Shaun Livingston	-1.3	92.5	630
Draymond Green	Klay Thompson	Festus Ezeli	-1.4	151.4	721
Stephen Curry	Klay Thompson	Festus Ezeli	-1.5	152.4	694
Draymond Green	Klay Thompson	Shaun Livingston	-1.6	160.5	929
Draymond Green	Stephen Curry	Brandon Rush	-1.7	153.9	1116
Draymond Green	Andre Iguodala	Andrew Bogut	-2.4	79.4	535
Stephen Curry	Andre Iguodala	Andrew Bogut	-2.8	70.2	541
Draymond Green	Stephen Curry	Leandro Barbosa	-3.0	126.9	687
Stephen Curry	Klay Thompson	Shaun Livingston	-3.0	121.3	530

Table 22: Worst triples for GSW with at least 500 possessions.

Player	SCLP	PM	Poss
Isaiah Thomas	3.4	236.5	5388
Avery Bradley	3.3	228.5	5099
Jae Crowder	3.1	219.5	4685
Jared Sullinger	3	210.8	3828
Amir Johnson	2	172.8	3580
Kelly Olynyk	1.7	154.8	2835
Marcus Smart	0.9	125.3	3407
Evan Turner	-0.2	81.1	4577
Jonas Jerebko	-0.4	71.5	2346
RJ Hunter	-3.3	-18.3	624
Jordan Mickey	-3.6	6	106
David Lee	-3.6	-35	945
Tyler Zeller	-3.7	-44.3	1442
James Young	-4	-31.3	392
Terry Rozier	-4.1	-43	616

Table 23: First order effects for BOS.

P1	P2	SCLP	PM	Poss
Isaiah Thomas	Avery Bradley	3.5	229.8	3564
Evan Turner	Jonas Jerebko	3.0	109.3	1945
Marcus Smart	Kelly Olynyk	3.0	141.8	1298
Avery Bradley	Jared Sullinger	2.7	191.2	2969
Tyler Zeller	RJ Hunter	2.6	8.5	261
Isaiah Thomas	Jared Sullinger	2.5	188.6	3315
Isaiah Thomas	Jae Crowder	2.3	187.0	3668
Jae Crowder	Amir Johnson	2.3	162.3	2594
Isaiah Thomas	Amir Johnson	2.1	165.8	3175
Kelly Olynyk	Jonas Jerebko	2.0	95.6	1030
Isaiah Thomas	Evan Turner	-2.2	1.0	2462
Avery Bradley	Tyler Zeller	-2.5	-38.0	674
Jared Sullinger	Jonas Jerebko	-2.5	-2.4	386
Isaiah Thomas	Tyler Zeller	-2.9	-41.5	455
Avery Bradley	Terry Rozier	-3.4	-41.9	160

Table 24: Top ten and bottom five second order effects for BOS with at least 150 possessions.

P1	P2	P3	SCLP	PM	Poss
Evan Turner	Kelly Olynyk	Jonas Jerebko	2.9	110.1	879
Isaiah Thomas	Avery Bradley	Jared Sullinger	2.7	177.7	2642
Avery Bradley	Jae Crowder	Jared Sullinger	2.3	139.3	2216
Isaiah Thomas	Avery Bradley	Jae Crowder	2.2	154.8	2572
Isaiah Thomas	Avery Bradley	Amir Johnson	2.0	137.5	2351
Evan Turner	Marcus Smart	Jonas Jerebko	2.0	93.7	1159
Jae Crowder	Evan Turner	Jonas Jerebko	2.0	61.2	460
Isaiah Thomas	Jae Crowder	Jared Sullinger	1.9	140.7	2533
Isaiah Thomas	Marcus Smart	Kelly Olynyk	1.8	85.5	464
Avery Bradley	Jae Crowder	Amir Johnson	1.7	107.3	1894
Avery Bradley	Jae Crowder	Evan Turner	-1.8	-7.9	708
Isaiah Thomas	Evan Turner	Tyler Zeller	-1.8	-68.4	305
Isaiah Thomas	Evan Turner	Kelly Olynyk	-1.8	-30.9	870
Avery Bradley	Jared Sullinger	Jonas Jerebko	-2.3	-11.7	194
Isaiah Thomas	Avery Bradley	Jonas Jerebko	-2.4	-1.6	290

Table 25: Top ten and bottom five third order effects for BOS with at least 150 possessions.

P1	P2	P3	P4	SCLP	PM	Poss
Avery Bradley	Evan Turner	Kelly Olynyk	Jonas Jerebko	3.1	71.8	375
Evan Turner	Marcus Smart	Kelly Olynyk	Jonas Jerebko	2.7	88.0	526
Isaiah Thomas	Avery Bradley	Jae Crowder	Jared Sullinger	2.6	120.0	2014
Isaiah Thomas	Avery Bradley	Evan Turner	Jared Sullinger	2.4	76.8	584
Avery Bradley	Evan Turner	Marcus Smart	Jonas Jerebko	2.1	62.8	526
Avery Bradley	Jae Crowder	Jared Sullinger	Kelly Olynyk	1.9	59.7	247
Avery Bradley	Marcus Smart	Kelly Olynyk	Jonas Jerebko	1.7	55.6	304
Isaiah Thomas	Avery Bradley	Jae Crowder	Kelly Olynyk	1.7	62.8	432
Avery Bradley	Evan Turner	Jared Sullinger	Amir Johnson	1.6	42.2	343
Avery Bradley	Evan Turner	Marcus Smart	Kelly Olynyk	1.6	44.3	423
Jae Crowder	Evan Turner	Jared Sullinger	Marcus Smart	-1.0	-21.2	180
Isaiah Thomas	Avery Bradley	Jae Crowder	Marcus Smart	-1.1	2.1	281
Evan Turner	Marcus Smart	Jonas Jerebko	Tyler Zeller	-1.1	1.5	408
Isaiah Thomas	Avery Bradley	Amir Johnson	Marcus Smart	-1.3	4.6	322
Isaiah Thomas	Avery Bradley	Evan Turner	Kelly Olynyk	-2.6	-24.2	225

Table 26: Top ten and bottom five fourth order effects for BOS with at least 150 possessions.

P1	P2	P3	P4	P5	SCLP	PM	Poss
Avery Bradley	Evan Turner	Marcus Smart	Kelly Olynyk	Jonas Jerebko	6.1	63.0	257
Isaiah Thomas	Avery Bradley	Jae Crowder	Jared Sullinger	Kelly Olynyk	3.4	41.9	202
Isaiah Thomas	Avery Bradley	Jae Crowder	Jared Sullinger	Amir Johnson	2.9	48.8	1413
Isaiah Thomas	Avery Bradley	Evan Turner	Jared Sullinger	Amir Johnson	2.4	33.8	256
Isaiah Thomas	Avery Bradley	Jae Crowder	Evan Turner	Jared Sullinger	1.1	23.7	148
Isaiah Thomas	Avery Bradley	Jae Crowder	Amir Johnson	Kelly Olynyk	-1.2	7.3	107
Isaiah Thomas	Avery Bradley	Jared Sullinger	Amir Johnson	Marcus Smart	-1.6	-3.8	128
Isaiah Thomas	Jae Crowder	Evan Turner	Jared Sullinger	Amir Johnson	-1.8	-7.0	105

Table 27: Top five and bottom three fifth order effects for BOS with at least 100 possessions.